Beyond Catoni: Sharper Rates for Heavy-Tailed and Robust Mean Estimation

Shivam Gupta (UT Austin), Samuel B. Hopkins (MIT), Eric Price (UT Austin)

One-Dimensional Mean Estimation

 Given n samples x₁,..., x_n from a variance σ² distribution, would like to produce an estimate of the mean μ

One-Dimensional Mean Estimation

- Given n samples x₁,..., x_n from a variance σ² distribution, would like to produce an estimate of the mean μ
- When the distribution is **Gaussian**, the empirical mean is within $\sigma \sqrt{\frac{2\log \frac{1}{\delta}}{n}}$ of the true mean μ with probability 1δ . (Chernoff Bound)

One-Dimensional Mean Estimation

- Given n samples x₁,..., x_n from a variance σ² distribution, would like to produce an estimate of the mean μ
- When the distribution is **Gaussian**, the empirical mean is within $\sigma \sqrt{\frac{2\log \frac{1}{\delta}}{n}}$ of the true mean μ with probability 1δ . (Chernoff Bound)
- For the general case:

Estimator	Error
Empirical Mean	$\sigma \sqrt{\frac{1}{n\delta}}$
Median-of-means	$19.2 \cdot \sigma \sqrt{\frac{\log \frac{1}{\delta}}{n}}$
Catoni (2012)	$\sigma \sqrt{\frac{2\log \frac{1}{\delta}}{n}} \cdot (1 + o(1))$

Table: One-dimensional Estimators

In *d*-dimensional estimation, we are given iid samples x₁,..., x_n ∈ ℝ^d, with Cov(x_i) ≤ σ²I_d and want to compute an estimate of the mean μ.

- In *d*-dimensional estimation, we are given iid samples x₁,..., x_n ∈ ℝ^d, with Cov(x_i) ≤ σ²I_d and want to compute an estimate of the mean μ.
- For simplicity, we will focus on the $\sigma = 1$ case.

- In *d*-dimensional estimation, we are given iid samples x₁,..., x_n ∈ ℝ^d, with Cov(x_i) ≤ σ²I_d and want to compute an estimate of the mean μ.
- For simplicity, we will focus on the $\sigma = 1$ case.

Estimator	Error	Notes
Empirical Mean	$\sqrt{\frac{d}{n}} + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}$	Gaussian/Light-tailed distributions
Catoni (2012) + Net	$\sqrt{\frac{2d}{d+1}} \cdot \left(O\left(\sqrt{\frac{d}{n}}\right) + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}\right)$	Any distribution
Catoni (2012) + PAC-Bayes	$\sqrt{\frac{2d}{d+1}} \cdot \left(\sqrt{\frac{d}{n}} + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}\right)$	Any distribution
Lee, Valiant (2022)	$\sqrt{\frac{d}{n}}$	Any distribution, when $d \gg \log^2 rac{1}{\delta}$

Table: Prior Estimators

- In *d*-dimensional estimation, we are given iid samples x₁,..., x_n ∈ ℝ^d, with Cov(x_i) ≤ σ²I_d and want to compute an estimate of the mean μ.
- For simplicity, we will focus on the $\sigma = 1$ case.

Estimator	Error	Notes
Empirical Mean	$\sqrt{\frac{d}{n}} + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}$	Gaussian/Light-tailed distributions
Catoni (2012) + Net	$\sqrt{\frac{2d}{d+1}} \cdot \left(O\left(\sqrt{\frac{d}{n}}\right) + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}\right)$	Any distribution
Catoni (2012) + PAC-Bayes	$\sqrt{\frac{2d}{d+1}} \cdot \left(\sqrt{\frac{d}{n}} + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}\right)$	Any distribution
Lee, Valiant (2022)	$\sqrt{\frac{d}{n}}$	Any distribution, when $d \gg \log^2 rac{1}{\delta}$

Table: Prior Estimators

• When $\log \frac{1}{\delta} \gg d$, is the $\sqrt{\frac{2d}{d+1}}$ -factor loss over the Gaussian rate necessary?

- In *d*-dimensional estimation, we are given iid samples x₁,..., x_n ∈ ℝ^d, with Cov(x_i) ≤ σ²I_d and want to compute an estimate of the mean μ.
- For simplicity, we will focus on the $\sigma = 1$ case.

Estimator	Error	Notes
Empirical Mean	$\sqrt{\frac{d}{n}} + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}$	Gaussian/Light-tailed distributions
Catoni (2012) + Net	$\sqrt{\frac{2d}{d+1}} \cdot \left(O\left(\sqrt{\frac{d}{n}}\right) + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}\right)$	Any distribution
Catoni (2012) + PAC-Bayes	$\sqrt{\frac{2d}{d+1}} \cdot \left(\sqrt{\frac{d}{n}} + \sqrt{\frac{2\log\frac{1}{\delta}}{n}}\right)$	Any distribution
Lee, Valiant (2022)	$\sqrt{\frac{d}{n}}$	Any distribution, when $d \gg \log^2 rac{1}{\delta}$

Table: Prior Estimators

- When log $\frac{1}{\delta} \gg d$, is the $\sqrt{\frac{2d}{d+1}}$ -factor loss over the Gaussian rate necessary?
- We show that the answer is **no** we show an estimator with error $(1 \tau) \cdot \sqrt{\frac{2d}{d+1}} \cdot \left(\sqrt{\frac{2\log \frac{1}{\delta}}{n}}\right)$ for a small constant $\tau > 0$.







• Suppose, we have an estimate $\hat{\mu}_{\mathbf{v}} = \langle \mu, \mathbf{v} \rangle \pm \epsilon$ for every unit vector \mathbf{v} .



• Taking the center of the enclosing sphere of these confidence regions is guaranteed to be within $\sqrt{\frac{2d}{d+1}} \cdot \epsilon$ of μ in ℓ_2 (Jung's theorem).



• Using a net argument with Catoni's estimate in each direction v, combined with this argument produces $\hat{\mu}$ with

$$\|\widehat{\mu}-\mu\| \leq \sqrt{rac{2d}{d+1}} \cdot \sigma \left(O\left(\sqrt{rac{d}{n}}
ight) + \sqrt{rac{2\lograc{1}{\delta}}{n}}
ight)$$

5/10

1. Let μ_0 be estimate via Median-of-means on fraction of the *n* samples

1. Let μ_0 be estimate via Median-of-means on fraction of the *n* samples 2. With $T = \sqrt{\frac{n}{2 \log \frac{1}{\delta}}}$, refine the estimate:

$$\widehat{\mu} = \mu_0 + \frac{1}{n} \sum_{i=1}^n T\psi\left(\frac{x_i - \mu_0}{T}\right)$$

1. Let μ_0 be estimate via Median-of-means on fraction of the *n* samples 2. With $T = \sqrt{\frac{n}{2 \log \frac{1}{\delta}}}$, refine the estimate:

$$\widehat{\mu} = \mu_0 + \frac{1}{n} \sum_{i=1}^n T\psi\left(\frac{x_i - \mu_0}{T}\right)$$

• When $\psi(x) = x$, then the estimate $\widehat{\mu}$ is just the empirical mean.

1. Let μ_0 be estimate via Median-of-means on fraction of the *n* samples 2. With $T = \sqrt{\frac{n}{2 \log \frac{1}{2}}}$, refine the estimate:

$$\widehat{\mu} = \mu_0 + \frac{1}{n} \sum_{i=1}^n T\psi\left(\frac{x_i - \mu_0}{T}\right)$$

- When $\psi(x) = x$, then the estimate $\widehat{\mu}$ is just the empirical mean.
- Catoni prescribes a specific way of selecting ψ to downweight outliers that achieves the optimal $\sigma \sqrt{\frac{2\log \frac{1}{\delta}}{n}}$ error.

1. Let μ_0 be estimate via Median-of-means on fraction of the *n* samples 2. With $T = \sqrt{\frac{n}{2 \log \frac{1}{2}}}$, refine the estimate:

$$\widehat{\mu} = \mu_0 + \frac{1}{n} \sum_{i=1}^n T\psi\left(\frac{x_i - \mu_0}{T}\right)$$

- When $\psi(x) = x$, then the estimate $\widehat{\mu}$ is just the empirical mean.
- Catoni prescribes a specific way of selecting ψ to downweight outliers that achieves the optimal $\sigma \sqrt{\frac{2\log \frac{1}{\delta}}{n}}$ error.
- ψ satisfies: $-\log\left(1-x+\frac{x^2}{2}\right) \le \psi(x) \le \log\left(1+x+\frac{x^2}{2}\right)$



• Estimate is given by

$$\widehat{\mu} = \mu_0 + r(\mu_0) := \mu_0 + \frac{1}{n} \sum_{i=1}^n T\psi\left(\frac{x_i - \mu_0}{T}\right)$$

• Estimate is given by

$$\widehat{\mu} = \mu_0 + r(\mu_0) := \mu_0 + \frac{1}{n} \sum_{i=1}^n T\psi\left(\frac{x_i - \mu_0}{T}\right)$$

• Claim: $r(\mu_0)$ is $(\mu - \mu_0) \pm \sigma \sqrt{\frac{2 \log \frac{1}{\delta}}{n}}$ with probability $1 - \delta$.

• Estimate is given by

$$\widehat{\mu} = \mu_0 + r(\mu_0) := \mu_0 + \frac{1}{n} \sum_{i=1}^n T\psi\left(\frac{x_i - \mu_0}{T}\right)$$

• Claim: $r(\mu_0)$ is $(\mu - \mu_0) \pm \sigma \sqrt{\frac{2 \log \frac{1}{\delta}}{n}}$ with probability $1 - \delta$.

• Proof: Its MGF is given by

$$\mathbb{E}\left[\exp\left(\frac{n}{T}r(\mu_{0})\right)\right] = \prod_{i=1}^{n} \mathbb{E}\left[\exp\left(\psi\left(\frac{x_{i}-\mu_{0}}{T}\right)\right)\right]$$
$$\leq \prod_{i=1}^{n} \mathbb{E}\left[\left(1+\frac{x_{i}-\mu_{0}}{T}+\frac{(x_{i}-\mu_{0})^{2}}{2T^{2}}\right)\right] \text{ since } \psi(x) \leq \log(1+x+x^{2}/2)$$
$$\leq \exp\left(\frac{n}{T}(\mu-\mu_{0})+\frac{n}{2T^{2}}\cdot\left[\sigma^{2}(1+o(1))\right]\right)$$



• Catoni uses

$$-\log\left(1-x+rac{x^2}{2}
ight) \leq \psi(x) \leq \log\left(1+x+rac{x^2}{2}
ight)$$



• Catoni uses

$$-\log\left(1-x+rac{x^2}{2}
ight) \leq \psi(x) \leq \log\left(1+x+rac{x^2}{2}
ight)$$

• There is slack in the choice for outliers



Catoni uses

$$-\log\left(1-x+rac{x^2}{2}
ight) \leq \psi(x) \leq \log\left(1+x+rac{x^2}{2}
ight)$$

- There is slack in the choice for outliers
- Taking advantage of the slack gives smaller error when the distribution has many outliers

• Recall:
$$T = \sqrt{\frac{n}{2\log \frac{1}{\delta}}}$$
 is the *scale* of the outliers

• **Strategy for two-dimensional estimator:** The distribution of *x_i* is *either* **outlier-heavy**, or **outlier-light**

• Recall:
$$T = \sqrt{\frac{n}{2 \log \frac{1}{\delta}}}$$
 is the *scale* of the outliers

- **Strategy for two-dimensional estimator:** The distribution of x_i is *either* **outlier-heavy**, or **outlier-light**
 - 1. **outlier-heavy**: Elements less than T/100 contribute less than 99% of the variance. Then, the Catoni estimate with our improved ψ constraint has a sharper error rate.

• Recall:
$$T = \sqrt{\frac{n}{2\log \frac{1}{\delta}}}$$
 is the *scale* of the outliers

- **Strategy for two-dimensional estimator:** The distribution of *x_i* is *either* **outlier-heavy**, or **outlier-light**
 - 1. **outlier-heavy**: Elements less than T/100 contribute less than 99% of the variance. Then, the Catoni estimate with our improved ψ constraint has a sharper error rate.
 - 2. **Outlier-light**: Elements more than T/100 contribute less than 1% of the variance. So, we can *trim* samples past this threshold, and compute an empirical mean.

• Recall:
$$T = \sqrt{\frac{n}{2\log \frac{1}{\delta}}}$$
 is the *scale* of the outliers

- **Strategy for two-dimensional estimator:** The distribution of *x_i* is *either* **outlier-heavy**, or **outlier-light**
 - 1. **outlier-heavy**: Elements less than T/100 contribute less than 99% of the variance. Then, the Catoni estimate with our improved ψ constraint has a sharper error rate.
 - 2. **Outlier-light**: Elements more than T/100 contribute less than 1% of the variance. So, we can *trim* samples past this threshold, and compute an empirical mean.
- In both cases, we achieve an improved rate. We can test which case we are in using a small fraction of samples.

• Recall:
$$T = \sqrt{\frac{n}{2\log \frac{1}{\delta}}}$$
 is the *scale* of the outliers

- **Strategy for two-dimensional estimator:** The distribution of *x_i* is *either* **outlier-heavy**, or **outlier-light**
 - 1. **outlier-heavy**: Elements less than T/100 contribute less than 99% of the variance. Then, the Catoni estimate with our improved ψ constraint has a sharper error rate.
 - 2. **Outlier-light**: Elements more than T/100 contribute less than 1% of the variance. So, we can *trim* samples past this threshold, and compute an empirical mean.
- In both cases, we achieve an improved rate. We can test which case we are in using a small fraction of samples.
- A generalization of Jung's theorem allows us to lift this estimator to *d* dimensions with an **improved rate**.

Main Theorems

Heavy-Tailed Estimation, Upper Bound

Suppose $\log \frac{1}{\delta} \gg d$. There is an algorithm that takes *n* samples in \mathbb{R}^d with covariance of each sample $\Sigma \preceq \sigma^2 I$, and outputs an estimate $\hat{\mu}$ with

$$\|\widehat{\mu} - \mu\| \leq (1 - au) \cdot \sqrt{rac{2d}{d+1}} \cdot \sigma \sqrt{rac{2\lograc{1}{\delta}}{n}}$$

with $1 - \delta$ probability, for some constant $\tau > 10^{-6}$.

Main Theorems

Heavy-Tailed Estimation, Upper Bound

Suppose $\log \frac{1}{\delta} \gg d$. There is an algorithm that takes *n* samples in \mathbb{R}^d with covariance of each sample $\Sigma \preceq \sigma^2 I$, and outputs an estimate $\hat{\mu}$ with

$$\|\widehat{\mu} - \mu\| \leq (1 - au) \cdot \sqrt{rac{2d}{d+1}} \cdot \sigma \sqrt{rac{2\lograc{1}{\delta}}{n}}$$

with $1 - \delta$ probability, for some constant $\tau > 10^{-6}$.

Robust Mean Estimation, Lower Bound

For every $d \ge 1$ and $\varepsilon \le \frac{1}{2}$, every algorithm for robust estimation of d-dimensional distributions with covariance $\Sigma \preceq \sigma^2 I$ has error rate

$$\mathbb{E}[\|\widehat{\mu} - \mu\|] \geq \sqrt{\frac{2d}{d+1}} \cdot (1 + O(\varepsilon)) \cdot \sqrt{2\sigma^2 \varepsilon}$$

on some input distribution, in the population limit.

10/10